## Mathematical Statistics

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Lecture XIV, 7.06.2021

#### **BAYESIAN STATISTICS**

Some concluding remarks

### **Plan for Today**

- **1.** Bayesian Statistics
  - a priori and a posteriori distributions
  - Bayesian estimation:
    - Maximum a posteriori probability (MAP)
    - Bayes Estimator
- 2. Caution!



# Frequentist: unknown parameters are given (fixed), observed data are random

# Bayesian: observed data are given (fixed), parameters are random



Warsaw University Faculty of Economic Sciences Our knowledge about the unknown parameters is described by means of probability distributions, and additional knowledge may affect our description.

Knowledge:

- general
- specific

### Example: coin toss



- $\Box X_1, ..., X_n$  come from distribution  $P_{\theta}$ , with density  $f_{\theta}(\mathbf{x})$  conditional density given a specific value of  $\theta$  (likelihood function).
- $\square \ \mathscr{P} \text{family of probability distributions } P_{\theta}, \\ \text{indexed by the parameter } \theta \in \Theta$
- General knowledge: distribution Π over the parameter space  $\Theta$ , given by  $\pi(\theta)$  the so-called **a priori/prior** distribution of  $\theta$ ,

 $\theta \sim \Pi$ 



### **Bayesian Model – cont.**

Additional knowledge (specific, contextual): based on observation. We have a joint distribution of observations and  $\theta$ .

 $f(x_1, x_2, \dots, x_n, \theta) = f(x_1, x_2, \dots, x_n | \theta) \pi(\theta)$ 

on this basis we can derive the conditional distribution of  $\theta$  (given the observed data)  $\pi(\theta|x_1,...,x_n) = \frac{f(x_1,...,x_n|\theta)\pi(\theta)}{m(x_1,...,x_n)},$ where  $m(x_1,...,x_n) = \int_{\Theta} f(x_1,...,x_n|\theta)\pi(\theta)d\theta$ is a marginal distribution for the obs.

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 $\pi(\theta|x_1,...,x_n)$  is called the **a posteriori**/ **posterior** distribution, denoted  $\Pi_x$ 

The posterior distribution reflects all knowledge: general (initial) and specific (based on the observed data).

Grounds for Bayesian inference and modeling



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### Prior and posterior distributions: examples

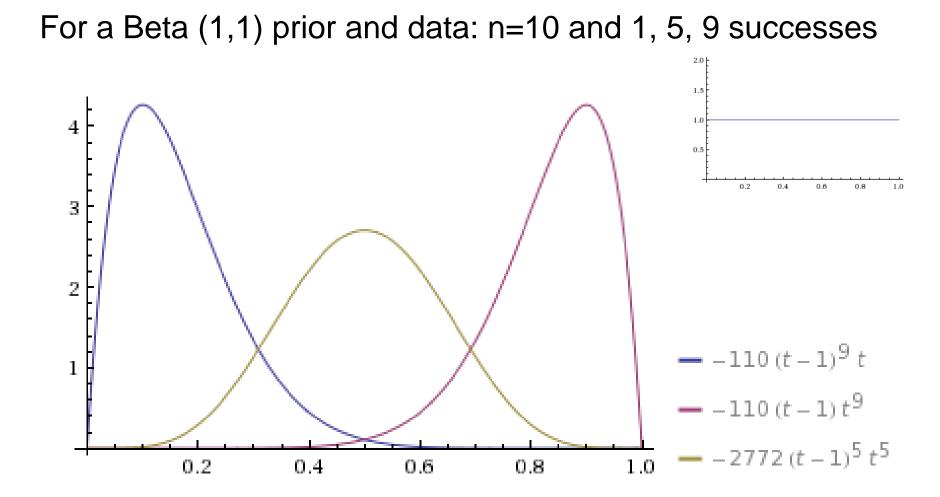
1.Let  $X_1, ..., X_n$  be IID r.v. from a 0-1 distr. with prob. of success  $\theta$ , let for  $\theta \in (0,1)$  $\pi(\theta) = \frac{\theta^{\alpha-1}(1-\theta)^{\beta-1}}{B(\alpha,\beta)}$ where  $B(\alpha,\beta) = \int_0^1 u^{\alpha-1}(1-u)^{\beta-1} du = \frac{\Gamma(\alpha)\Gamma(\beta)}{\Gamma(\alpha+\beta)}$ and  $\Gamma(\alpha) = \int_0^\infty u^{\alpha-1} \exp(-u) du = (\alpha-1)\Gamma(\alpha-1)$ Beta $(\alpha,\beta)$ distr with mean  $= \alpha/(\alpha+\beta)$ 

then the posterior distribution:

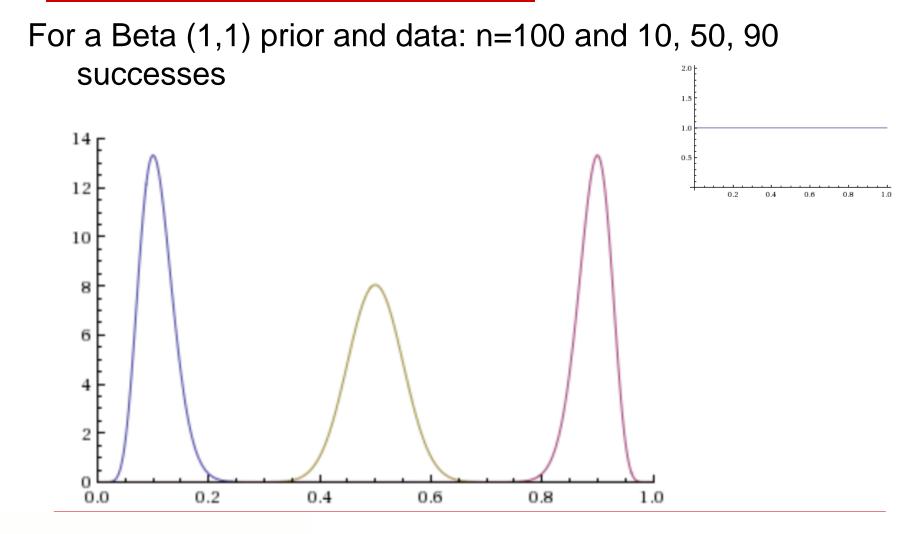
Beta
$$(\sum_{i=1}^{n} x_i + \alpha, n - \sum_{i=1}^{n} x_i + \beta)$$

**X** 

Warsaw University Faculty of Economic Sciences conjugate prior for Bernoulli distr.



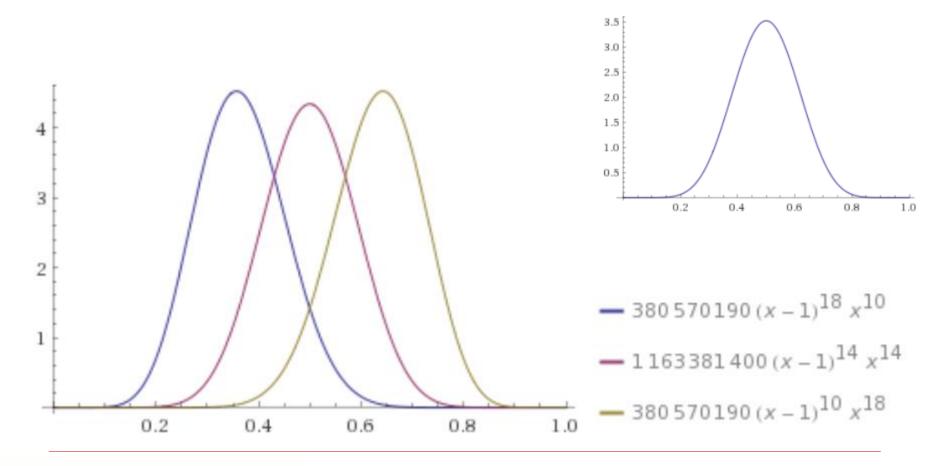




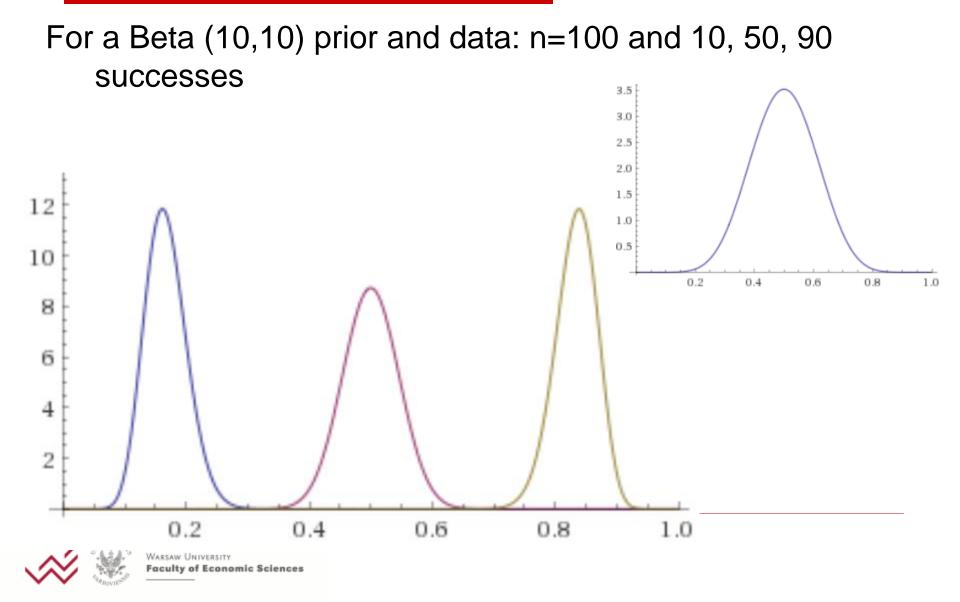


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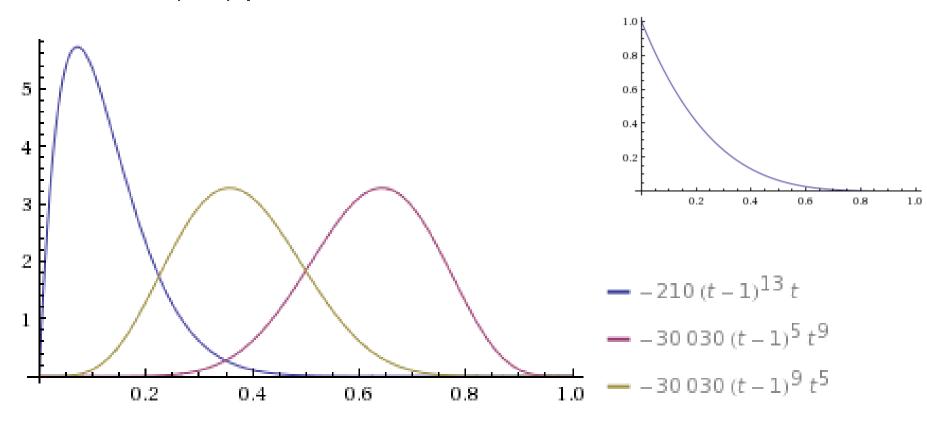




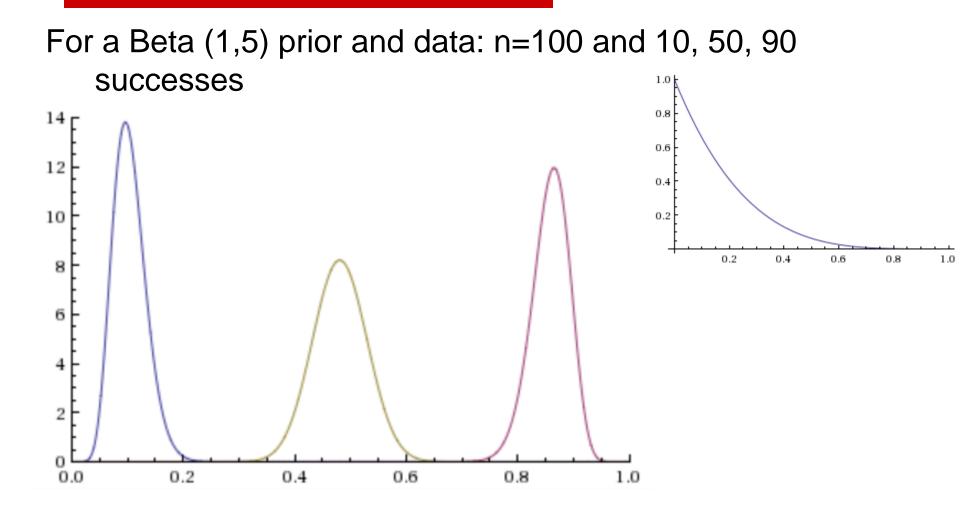










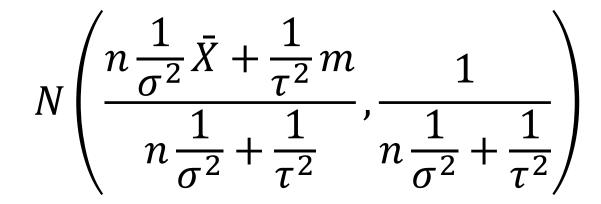




### Prior and posterior distributions: examples (2)

2. Let  $X_1, ..., X_n$  be IID r.v. from  $N(\theta, \sigma^2)$ , and  $\sigma^2$  known;  $\theta \sim N(m, \tau^2)$  for  $m, \tau$  known.

Then the posterior distribution for  $\theta$ .



conjugate prior for a normal distr.



Based on the Bayes approach, we can

- □ find estimates
- □ find an equivalent of confidence intervals
- verify hypotheses

make predictions



## Bayesian Most Probabale (BMP) / Maximum a posteriori Probability (MAP) estimate

Similar to ML estimation: the argument which maximizes the posterior distribution:

$$\pi(\hat{\theta}_{BMP}|x_1,\ldots,x_n) = \max_{\theta} \pi\left(\theta|x_1,\ldots,x_n\right)$$

i.e.

### $BMP(\theta) = \hat{\theta}_{BMP} = \operatorname{argmax}_{\theta} \pi \left( \theta | x_1, \dots, x_n \right)$



1. Let  $X_1, ..., X_n$  be IID r.v. from a Bernoulli distr. with prob. of success  $\theta$ ; for  $\theta \in (0,1)$ We know the posterior distribution:  $Beta(\sum_{i=1}^{n} x_i + \alpha, n - \sum_{i=1}^{n} x_i + \beta)$ we have max for  $BMP(\theta) = \frac{\sum_{i=1}^{n} x_i + \alpha - 1}{n + \beta + \alpha - 2}$ Beta( $\alpha, \beta$ ) distr; the mode of this distr for  $\alpha > 1, \beta > 1$ 

i.e. for 5 successes in 10 trials for a prior U(0,1) (i.e. Beta(1,1) distr.), we have  $BMP(\theta)=5/10=\frac{1}{2}$ 

and for 9 successes in 10 trials for the same prior distr., we have  $BMP(\theta)=9/10$ 



**2.** Let  $X_1, ..., X_n$  be IID r.v. from  $N(\theta, \sigma^2)$ , with  $\sigma^2$  known;  $\theta \sim N(m, \tau^2)$  for  $m, \tau$  known.

Then the posterior distr. for  $\theta$ :

$$BMP(\theta) = \frac{n\frac{1}{\sigma^2}\bar{X} + \frac{1}{\tau^2}m}{n\frac{1}{\sigma^2} + \frac{1}{\tau^2}}$$

$$N\left(\frac{n\frac{1}{\sigma^{2}}\bar{X} + \frac{1}{\tau^{2}}m}{n\frac{1}{\sigma^{2}} + \frac{1}{\tau^{2}}}, \frac{1}{n\frac{1}{\sigma^{2}} + \frac{1}{\tau^{2}}}\right)$$

i.e. if we have a sample of 5 obs 1.2; 1.7 ; 1.9 ; 2.1; 3.1 from distr. N( $\theta$ , 4) and the prior distr is  $\theta \sim N(1, 1)$ , then BMP( $\theta$ ) = (5 /4 \* 2 + 1)/(5/4 + 1) = 14/9  $\approx$  1.56

and if the prior distr were  $\theta \sim N(3, 1)$ , then

 $BMP(\theta) = (5/4 * 2 + 1*3)/(5/4 + 1) = 22/9 \approx 2.44$ 



SO

An estimation rule which minimizes the posterior expected value of a loss function

# $L(\theta, a) -$ **loss function**, depends on the true value of $\theta$ and the decision *a*.

e.g. if we want to estimate  $g(\theta)$ :  $L(\theta, a) = (g(\theta) - a)^2 - quadratic loss function$  $L(\theta, a) = |g(\theta) - a| - module loss function$ 



# We can also define the **accuracy of an estimate** for a given loss function :

 $acc(\Pi, \hat{g}(x)) = E(L(\theta, \hat{g}(x))|X = x) = \int_{\Theta} L(\theta, \hat{g}(x))\pi(\theta|x)d\theta$ 

(the average loss of the estimator for a given prior distribution and data, i.e. for a specific posterior distribution)



# The **Bayes Estimator** for a given loss function $L(\theta, a)$ is $\hat{g}_B$ such that

$$\forall x \quad acc(\Pi, \hat{g}_B(x)) = \min_a a \, cc(\Pi, a)$$

### For a quadratic loss function $(\theta - a)^2$ : $\hat{\theta}_B = E(\theta | X = x) = E(\Pi_x)$ For a module loss function $|\theta - a|^2$ : $\hat{\theta}_B = Med(\Pi_x)$



WARSAW UNIVERSITY Faculty of Economic Sciences more generally:  $E(g(\theta)|x)$ 

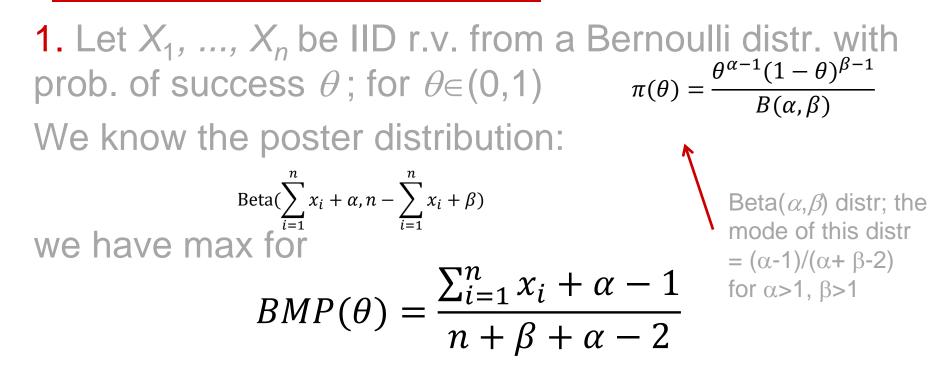
### **Bayes Estimator: Example (1)**

1. Let  $X_1, ..., X_n$  be IID r.v. from a Bernoulli distr. with prob. of success  $\theta$ ; for  $\theta \in (0,1)$ We know the posterior distribution:  $\underset{Beta(\sum_{i=1}^{n} x_i + \alpha, n - \sum_{i=1}^{n} x_i + \beta)}{\underset{\theta_B}{\text{Beta}(\alpha, \beta) \text{ distr with}}}$ Beta( $\alpha, \beta$ ) distr with mean  $= \alpha/(\alpha + \beta)$ 

i.e. for 5 successes in 10 trials for a prior U(0,1) (i.e. Beta(1,1) distr.), we have  $\hat{\theta}_B = 6/12 = \frac{1}{2}$ 

and for 9 successes in 10 trials for the same prior distr., we have  $\hat{\theta}_B = 10/12 = 5/6$ 





i.e. for 5 successes in 10 trials for a prior U(0,1) (i.e. Beta(1,1) distr.), we have  $BMP(\theta)=5/10 = \frac{1}{2}$ 

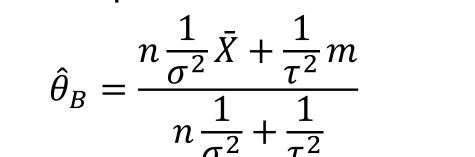
and for 9 successes in 10 trials for the same prior distr., we have  $BMP(\theta)=9/10$ 



### **Bayes Estimator: examples (2)**

**2.** Let  $X_1, ..., X_n$  be IID r.v. from  $N(\theta, \sigma^2)$ , with  $\sigma^2$  known;  $\theta \sim N(m, \tau^2)$  for  $m, \tau$  known.

Then the a posteriori distr for  $\theta$ :



$$N\left(\frac{n\frac{1}{\sigma^{2}}\bar{X} + \frac{1}{\tau^{2}}m}{n\frac{1}{\sigma^{2}} + \frac{1}{\tau^{2}}}, \frac{1}{n\frac{1}{\sigma^{2}} + \frac{1}{\tau^{2}}}\right)$$

i.e. if we have sa sample of 5 obs 1.2; 1.7 ; 1.9 ; 2.1; 3.1 from distr. N( $\theta$ , 4) and the prior distr is  $\theta \sim N(1, 1)$ , then

$$\hat{\theta}_B = (5/4 \times 2 + 1)/(5/4 + 1) = 14/9 \approx 1.56$$

and if the prior distr were  $\theta \sim N(3, 1)$ , then

$$\hat{P}_B = (5/4 \times 2 + 1 \times 3)/(5/4 + 1) = 22/9 \approx 2.44$$



SO

**2**. Let  $X_1, ..., X_n$  be IID r.v. from  $N(\theta, \sigma^2)$ , with  $\sigma^2$  known;  $\theta \sim N(m, \tau^2)$  for  $m, \tau$  known.

Then the a posteriori distr for  $\theta$ :

$$N\left(\frac{n\frac{1}{\sigma^{2}}\bar{X} + \frac{1}{\tau^{2}}m}{n\frac{1}{\sigma^{2}} + \frac{1}{\tau^{2}}}, \frac{1}{n\frac{1}{\sigma^{2}} + \frac{1}{\tau^{2}}}\right)$$

$$BMP(\theta) = \frac{n\frac{1}{\sigma^2}\bar{X} + \frac{1}{\tau^2}m}{n\frac{1}{\sigma^2} + \frac{1}{\tau^2}}$$

i.e. if we have sa sample of 5 obs 1.2; 1.7 ; 1.9 ; 2.1; 3.1 from distr. N( $\theta$ , 4) and the prior distr is  $\theta \sim N(1, 1)$ , then

$$BMP(\theta) = (5/4 \times 2 + 1)/(5/4 + 1) = 14/9 \approx 1.56$$

and if the prior distr were  $\theta \sim N(3, 1)$ , then

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 $BMP(\theta) = (5/4 * 2 + 1*3)/(5/4 + 1) = 22/9 \approx 2.44$ 



SO

### **Caution!**

- Tests should be designed BEFORE we start examining the data
- The only way to increase power and improve significance level simultaneously is by collecting more observations (frequently not possible if we work on existing data).
- **3**. Significant p-value does not mean effect is important/sizeable.



We examine if a new training has effect. The null hypothesis is that the training has no effect, and the alternative hypothesis is that it has effect. We use a 5% significance level for the test.

- A randomly selected school has completed this training, and after completion the statistical test returns a *P*-value equal to 4%.
- 25 different schools have completed this training. At one of the schools the test returned a *P*-value of 4%.

