Mathematical Statistics

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HYPOTHESIS TESTING III:

LR TEST FOR COMPOSITE HYPOTHESES

EXAMPLES OF ONE-SAMPLE TESTS

Plan for today

1. LR test for composite hypotheses
2. Examples of LR tests:
   - Model I: One- and two-sided tests for the mean in the normal model, σ^2 known
   - Model II: One- and two-sided tests for the mean in the normal model, σ^2 unknown

     + One- and two-sided tests for the variance
   - Model III: Tests for the mean, large samples
   - Model IV: Tests for the fraction, large samples

Testing simple hypotheses – reminder

We observe X. We want to test

H_0: $\theta = \theta_0$ against H_1: $\theta = \theta_1$.

(two simple hypotheses)

We can write it as:

H_0: $X \sim f_0$ against H_1: $X \sim f_1$,

where f_0 and f_1 are *densities* of distributions defined by θ_0 and θ_1 (i.e. P_0 and P_1)

Likelihood ratio test for simple hypotheses. Neyman-Pearson Lemma – reminder

H_0: $X \sim f_0$ against H_1: $X \sim f_1$

Let

$$C^* = \left\{ x \in \boldsymbol{X} : \frac{f_1(x)}{f_0(x)} > c \right\}$$

such that $P_0(C^*) = \alpha$ and $P_1(C^*) = 1 - \beta$

Then, for any $C \subseteq \mathcal{X}$:

if $P_0(C) \leq \alpha$, then $P_1(C) \leq 1 - \beta$.

(i.e.: the test with critical region C^* is the most powerful test for testing H_0 against H_1)

Neyman-Pearson Lemma – Example 1 reminder

Normal model: $X_1, X_2, \ldots, X_n$ are an IID sample from $N(\mu, \sigma^2)$, σ^2 is known

The most powerful test for

H_0: $\mu = 0$ against H_1: $\mu = 1$. ← $\mu_0 < \mu_1$

At significance level α :

$$C^* = \left\{(x_1, x_2, \ldots, x_n) : \bar{X} > {u_{1-\alpha}\sigma}/{\sqrt{n}}\right\}$$

If we had

H_0: $\mu = 0$ against H_1: $\mu = -1$, then

$$C_1^* = \left\{(x_1, x_2, \ldots, x_n) : \bar{X} < {u_{1-\alpha}\sigma}/{\sqrt{n}}\right\}$$

Neyman-Pearson Lemma – Example 1 cont.

Power of the test

$$P_1(C^*) = P\left(\bar{X} > {}^{1.645\sigma}\!/_{\sqrt{n}} \,|\, \mu = 1\right) = \ldots.$$

$$= 1 - \Phi\left(1.645 - {}^{\mu_1 \cdot \sqrt{n}}\!/_{\sigma}\right) \approx 0.91$$

If we change α, μ_1, n – the power of the test....

Neyman-Pearson Lemma: Generalization of example 1

The same test is UMP for H_1: $\mu > 0$ and for

$$H_0: \mu \leq 0 \text{ against } H_1: \mu > 0$$

more generally: under additional assumptions about the family of distributions, the same test is UMP for testing

$$H_0: \mu \leq \mu_0 \text{ against } H_1: \mu > \mu_0$$

Note the change of direction of the inequality in the condition when testing

$$H_0: \mu \geq \mu_0 \text{ against } H_1: \mu < \mu_0$$

Neyman-Pearson Lemma – Example 2

Exponential model: $X_1, X_2, \ldots, X_n$ are an IID sample from distr $\exp(\lambda)$, $n = 10$.

MP test for

$$H_0: \lambda = ½ \text{ against } H_1: \lambda = ¼.$$

At significance level $\alpha = 0.05$:

$$C^* = \left\{(x_1, x_2, \ldots, x_{10}) : \sum x_i > 31.41\right\}$$

E.g. for a sample: 2; 0.9; 1.7; 3.5; 1.9; 2.1; 3.7; 2.5; 3.4; 2.8:

$\Sigma = 24.5 \rightarrow$ no grounds for rejecting H_0.

$\exp(\lambda) = \Gamma(1, \lambda) \quad \Gamma(a, \lambda) + \Gamma(b, \lambda) = \Gamma(a + b, \lambda) \quad \Gamma({}^{n}/_{2}, {}^{1}/_{2}) = \chi^2(n)$

Neyman-Pearson Lemma – Example 2’

Exponential model: $X_1, X_2, \ldots, X_n$ are an IID sample from distr exp(λ), $n = 10$.

MP test for

H_0: $\lambda = ½$ against H_1: $\lambda = ¾$.

At significance level $\alpha = 0.05$:

$$C^* = \left\{(x_1, x_2, \ldots, x_{10}) : \sum x_i < 10.85\right\}$$

E.g. for a sample: 2; 0.9; 1.7; 3.5; 1.9; 2.1; 3.7; 2.5; 3.4; 2.8:

$\Sigma = 24.5 \rightarrow$ no grounds for rejecting H_0.

$$\exp(\lambda) = \Gamma(1, \lambda) \quad \Gamma(a, \lambda) + \Gamma(b, \lambda) = \Gamma(a + b, \lambda) \quad \Gamma({}^{n}/_{2}, {}^{1}/_{2}) = \chi^2(n)$$

Example 2 cont.

The test

$$C^* = \left\{ (x_1, x_2, \ldots, x_{10}) : \sum x_i > 31.41 \right\}$$

is UMP for H_0: $\lambda \geq$ ½ against H_1: $\lambda <$ ½

The test

$$C^* = \left\{ (x_1, x_2, \ldots, x_{10}) : \sum x_i < 10.85 \right\}$$

is UMP for H_0: $\lambda \leq$ ½ against H_1: $\lambda >$ ½

Likelihood ratio test for composite hypotheses

$X \sim P_\theta$, $\{P_\theta \colon \theta \in \Theta\}$ – family of distributions

We are testing H_0: $\theta \in \Theta_0$ against H_1: $\theta \in \Theta_1$ such that $\Theta_0 \cap \Theta_1 = \varnothing$, $\Theta_0 \cup \Theta_1 = \Theta$

Let

H_0: $X \sim f_0(\theta_0, \cdot)$ for some $\theta_0 \in \Theta_0$.

H_1: $X \sim f_1(\theta_1, \cdot)$ for some $\theta_1 \in \Theta_1$,

where f_0 and f_1 are densities (for $\theta \in \Theta_0$ and $\theta \in \Theta_1$, respectively)

Just like in the N-P Lemma, but models are statistic – contain unknow parameters. We proceed similarly.

Likelihood ratio test for composite hypotheses – cont.

Test statistic: $$\lambda = \frac{\sup_{\theta_1 \in \Theta_1} f_1(\theta_1, X)}{\sup_{\theta_0 \in \Theta_0} f_0(\theta_0, X)}$$

or $$\lambda = \frac{f_1(\hat{\theta}_1, X)}{f_0(\hat{\theta}_0, X)}$$

where $\hat{\theta}_0, \hat{\theta}_1$ are MLE for the null and alternative hypothesis models

We reject H_0 if $\lambda > c$ for a constant c (determined according to significance level)

Likelihood ratio test for composite hypotheses – justification

Just like in the Neyman-Pearson Lemma, we compare the "highest chance of obtaining observation *X*, when the alternative is true" to the "highest chance of obtaining observation *X*, when the null is true"; we reject the null hypothesis in favor of the alternative if this ratio is very unfavorable for the null.

Likelihood ratio test for composite hypotheses – alternative version

Test statistic: $$\tilde{\lambda} = \frac{\sup_{\theta \in \Theta} f\,(\theta, X)}{\sup_{\theta_0 \in \Theta_0} f_0\,(\theta_0, X)}$$

or $$\tilde{\lambda} = \frac{f\,(\hat{\theta}, X)}{f_0(\hat{\theta}_0, X)}$$

where $\hat{\theta}, \hat{\theta}_0$ are the ML estimators for the model without restrictions and for the null model, respectively.

We reject H_0 if $\tilde{\lambda} > \tilde{c}$ for a constant $\tilde{c}$.

more convenient if the null is simple or if models are nested

Likelihood ratio test for composite hypotheses – properties

For some models with composite hypotheses the UMPT *does not exist* (so the LR test will not be UMP because there is no such test)

e.g. testing H_0: $\theta = \theta_0$ against H_1: $\theta \neq \theta_0$ if the family of distributions has a *monotonic LR property*, i.e. $f_1(x)/f_0(x)$ is an increasing function of a statistic $T(x)$ for any f_0 and f_1 corresponding to parameters $\theta_0 < \theta_1$.

In order to have UMPT for H_0: $\theta = \theta_0$ against H_1: $\theta > \theta_0$ we would need a critical region of the type $T(x)>c$, and to have a UMPT for H_0: $\theta = \theta_0$ against H_1: $\theta < \theta_0$ we would need a critical region of the type $T(x)<c$, so it is impossible to find a UMPT for H_1: $\theta \neq \theta_0$.

Likelihood ratio test: special cases

The exact form of the test depends on the distribution.

In many cases, finding the distribution is hard/complicated (in many such cases, we use the asymptotic properties of the LR test instead of precise formulae)

Notation

$x_{something}$ **always** means a quantile of rank something

Model I: comparing the mean

Normal model: $X_1, X_2, \ldots, X_n$ are an IID sample from $N(\mu, \sigma^2)$, where σ^2 is **known**

H_0: $\mu = \mu_0$

Test statistic:

$$U = \frac{\bar{X} - \mu_0}{\sigma}\sqrt{n} \sim N(0,1)$$

H_0: $\mu = \mu_0$ against H_1: $\mu > \mu_0$

critical region $C^* = \{x : U(x) > u_{1-\alpha}\}$

H_0: $\mu = \mu_0$ against H_1: $\mu < \mu_0$

critical region $C^* = \{x : U(x) < u_\alpha = -u_{1-\alpha}\}$

H_0: $\mu = \mu_0$ against H_1: $\mu \neq \mu_0$

critical region $C^* = \{x : |U(x)| > u_{1-\alpha/2}\}$

Model I: example

Let $X_1, X_2, \ldots, X_{10}$ be an IID sample from $N(\mu, 1^2)$:

-1.21 -1.37 0.51 0.37 -0.75 0.44 1.20 -0.96 -1.14 -1.40

Is $\mu = 0$? (for $\alpha = 0.05$)

In the sample: mean = -0.43, ~~variance = 0.92~~

Test statistic:

$$U = \frac{-0.43 - 0}{1}\sqrt{10} \approx -1.36$$

H_0: $\mu = 0$ against H_1: $\mu \neq 0$, $u_{0.975} \approx 1.96$ (p-value ≈ 0.172)

H_0: $\mu = 0$ against H_1: $\mu < 0$, $u_{0.05} \approx -1.64$ (p-value ≈ 0.086)

H_0: $\mu = 0$ against H_1: $\mu > 0$, $u_{0.95} \approx 1.64$ (p-value ≈ 0.914)

→ in none of these cases are there grounds to reject H_0 for $\alpha = 0.05$

→but we would reject H_0: $\mu = 0$ in favor of H_1: $\mu < 0$ for $\alpha = 0.1$

Model II: comparing the mean

Normal model: $X_1, X_2, \ldots, X_n$ are an IID sample from $N(\mu, \sigma^2)$, where σ^2 is **unknown**

H_0: $\mu = \mu_0$

Test statistic:

$$T = \frac{\bar{X} - \mu_0}{S}\sqrt{n} \sim t\,(n-1)$$

H_0: $\mu = \mu_0$ against H_1: $\mu > \mu_0$

critical region $C^* = \{x : T\,(x) > t_{1-\alpha}(n-1)\}$

H_0: $\mu = \mu_0$ against H_1: $\mu < \mu_0$

critical region $C^* = \{x : T\,(x) < t_{\alpha}(n-1)\}$

H_0: $\mu = \mu_0$ against H_1: $\mu \neq \mu_0$

critical region $C^* = \{x : |T(x)| > t_{1-\alpha/2}(n-1)\}$

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$$t_{\alpha}(n-1) = -t_{1-\alpha}(n-1)$$

Model II: example (mean)

Let $X_1, X_2, \ldots, X_{10}$ be an IID sample from $N(\mu, \sigma^2)$:

-1.21 -1.37 0.51 0.37 -0.75 0.44 1.20 -0.96 -1.14 -1.40

Is $\mu = 0$? (for $\alpha = 0.05$)

In the sample: mean = -0.43, variance = 0.92

Test statistic:

$$U = \frac{-0.43 - 0}{\sqrt{0.92}}\sqrt{10} \approx -1.42$$

H_0: $\mu = 0$ vs H_1: $\mu \neq 0$, $t_{0.975}(9) \approx 2.26$ (p-value $\approx$ 0.188)

H_0: $\mu = 0$ vs H_1: $\mu < 0$, $t_{0.05}(9) \approx -1.83$ (p-value $\approx$ 0.094)

H_0: $\mu = 0$ vs H_1: $\mu > 0$, $t_{0.95}(9) \approx 1.83$ (p-value $\approx$ 0.906)

→ in none of these cases are there grounds to reject H_0 for $\alpha = 0.05$

→but we would reject H_0: $\mu = 0$ in favor of H_1: $\mu < 0$ for $\alpha = 0.1$

Model II: comparing the variance

Normal model: $X_1, X_2, \ldots, X_n$ are an IID sample from $N(\mu, \sigma^2)$, where σ^2 is **unknown**

H_0: $\sigma = \sigma_0$

Test statistic:

$$\chi^2 = \frac{(n-1)S^2}{\sigma_0^2} \sim \chi^2(n-1)$$

H_0: $\sigma = \sigma_0$ against H_1: $\sigma > \sigma_0$

critical region $C^* = \{x : \chi^2(x) > \chi^2_{1-\alpha}(n-1)\}$

H_0: $\sigma = \sigma_0$ against H_1: $\sigma < \sigma_0$

critical region $C^* = \{x : \chi^2(x) < \chi^2_{\alpha}(n-1)\}$

H_0: $\sigma = \sigma_0$ against H_1: $\sigma \neq \sigma_0$

critical region $C^* = \{x : \chi^2(x) < \chi^2_{\alpha/2}(n-1) \vee \chi^2(x) > \chi^2_{1-\alpha/2}(n-1)\}$

Model II: example (variance)

Let $X_1, X_2, \dots, X_{10}$ be an IID sample from $N(\mu, \sigma^2)$:

-1.21 -1.37 0.51 0.37 -0.75 0.44 1.20 -0.96 -1.14 -1.40

Is $\sigma = 1$? (for $\alpha = 0.05$)

In the sample: variance = 0.92

Test statistic:

$$\chi^2 = \frac{9 \cdot 0.92}{1} \approx 8.28$$

H_0: $\sigma = 1$ against H_1: $\sigma > 1$ $\quad \chi^2_{0.95} \approx 16.92$

H_0: $\sigma = 1$ against H_1: $\sigma < 1$ $\quad \chi^2_{0.05} \approx 3.33$

H_0: $\sigma = 1$ against H_1: $\sigma \neq 1$ $\quad \chi^2_{0.025} \approx 2.70; \chi^2_{0.975} \approx 19.02$

$\rightarrow$ in none of these cases are there grounds to reject H_0 (for $\alpha = 0.05$)

Model III: comparing the mean

Asymptotic model: $X_1, X_2, \ldots, X_n$ are an IID sample from a distribution with mean μ and variance (unknown), n – large.

H_0: $\mu = \mu_0$

Test statistic:

$$T = \frac{\bar{X} - \mu_0}{S}\sqrt{n}$$

has, for large *n, an approximate* distribution N(0,1)

H_0: $\mu = \mu_0$ against H_1: $\mu > \mu_0$

critical region $C^* = \{x : T(x) > u_{1-\alpha}\}$

H_0: $\mu = \mu_0$ against H_1: $\mu < \mu_0$

critical region $C^* = \{x : T(x) < u_{\alpha} = -u_{1-\alpha}\}$

H_0: $\mu = \mu_0$ against H_1: $\mu \neq \mu_0$

critical region $C^* = \{x : |T(x)| > u_{1-\alpha/2}\}$

Model IV: comparing the fraction

Asymptotic model: $X_1, X_2, \ldots, X_n$ are an IID sample from a two-point distribution, n – large.

$$P_p(X = 1) = p = 1 - P_p(X = 0)$$

H_0: $p = p_0$

Test statistic:

$$U^* = \frac{\bar{X} - p_0}{\sqrt{p_0(1 - p_0)}}\sqrt{n} = \frac{\hat{p} - p_0}{\sqrt{p_0(1 - p_0)}}\sqrt{n}$$

has *an approximate* distribution N(0,1) for large n

H_0: $p = p_0$ against H_1: $p > p_0$

critical region $C^* = \{x : U^*(x) > u_{1-\alpha}\}$

H_0: $p = p_0$ against H_1: $p < p_0$

critical region $C^* = \{x : U^*(x) < u_\alpha = -u_{1-\alpha}\}$

H_0: $p = p_0$ against H_1: $p \neq p_0$

critical region $C^* = \{x : |U^*(x)| > u_{1-\alpha/2}\}$

Model IV: example

We toss a coin 400 times. We get 180 heads. Is the coin symmetric?

H_0: $p = ½$

$$U^* = \frac{(180/400 - 1/2)}{\sqrt{1/2(1-1/2)}}\sqrt{400} = -2$$

for $\alpha = 0.05$ and H_1: $p \neq ½$ we have $u_{0.975} = 1.96 \rightarrow$ we reject H_0

for $\alpha = 0.05$ and H_1: $p < ½$ we have $u_{0.05} = -u_{0.95} = -1.64$

$\rightarrow$ we reject H_0

for $\alpha = 0.01$ and H_1: $p \neq ½$ we have $u_{0.995} = 2.58$

$\rightarrow$ we do not reject H_0

for $\alpha = 0.01$ and H_1: $p < ½$ we have $u_{0.01} = -u_{0.99} = -2.33$

$\rightarrow$ we do not reject H_0

p-value for H_1: $p \neq ½$: 0.044 **p-value for H_1: $p < ½$: 0.022**